

What is this text about?

How to use Long & Freese routines in multinomial logit problems.
Demonstrated with the example of chapter 6 of Long & Freese's textbook¹.
This text is a summarization of the Long & Freese chapter.

Whom is this text for?

Who urgently need these routines but do not have the textbook or time to read it.

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¹ Long JS, Freese J (2006) Regression for Categorical Dependent Variables Using Stata. Stata Press

Reinstall SPOST

You might have to reinstall the Long & Freese routines SPOST.
The problems for STATA 10 are now fixed.

See: http://www.indiana.edu/~jslsoc/web_spost/sp_news.htm

How to reinstall SPOST (for STATA 9 and STATA 10):

net from <http://www.indiana.edu/~jslsoc/stata/>
net install spost9_ado, replace

The model ²

$$\pi_j(x) := P(Y = j | x)$$

$$\text{Baseline - category logit : } \left(\frac{\pi_j(x)}{\pi_J(x)} \right) \text{ for } j = 1, \dots, J$$

The multinomial logit model (MNL) = polytomous logit model = discrete choice model is:

$$\ln \left(\frac{\pi_j(x)}{\pi_J(x)} \right) = \alpha_j + \beta_j'x \quad \text{for } j = 1, \dots, J-1$$

where **J is the baseline category**.

The model simultaneously describes the effect of x on J-1 baseline category logits

Other pairs of response categories can be derived since:

$$\ln \left(\frac{\pi_a(x)}{\pi_b(x)} \right) = \ln \left(\frac{\pi_a(x)}{\pi_J(x)} \right) - \ln \left(\frac{\pi_b(x)}{\pi_J(x)} \right)$$

Response probabilities are derived as

$$\pi_j(x) = \frac{\exp(\alpha_j + \beta_j'x)}{1 + \sum_{h=1}^{J-1} \exp(\alpha_h + \beta_h'x)} = \frac{\exp(y_j)}{1 + \sum_{h=1}^{J-1} \exp(y_h)}$$

The odds ratio of outcome Y=j versus outcome Y=J for x=a versus x=b is

$$\text{OR}_{j,J}(a,b) = \frac{\frac{P(Y = j | a)}{P(Y = J | a)}}{\frac{P(Y = j | b)}{P(Y = J | b)}} = \frac{\frac{\pi_j(a)}{\pi_J(a)}}{\frac{\pi_j(b)}{\pi_J(b)}}$$

² Agresti A (2002) Categorical data analysis. Wiley, New York, pp. 268

The example file in Long & Freese consists of 337 cases from the 1982 General Social Survey (GSS)
<http://www.stata-press.com/data/lf2/nomocc2>

Variable	Label	Value labels
occ	Occupation	1=Menial 2=BlueCol 3=Craft 4=WhiteCol 5=Prof
white	Race: 1=white 0=nonwhite	
ed	Years of education	
exper	Years of work experience	

Input data file

```
open the file  
use http://www.stata-press.com/data/lf2/nomocc2,clear
```

do some labeling

```
label def lbwhite 0 NonWhite 1 White  
label val white lbwhite  
label variable white Race
```

creating a new variable occ3 which has 3 categories of types of work instead of the original 5 categories for variable occ:

```
recode occ (1/3=1 "Manual" ) (4=2 "WhiteCol") (5=3 "Prof"), gen(occ3) label (lbocc3) test  
label variable occ3 "Occupation simplified into 3 classes"
```

save the file

```
save "nomocc2_own.dta"
```

Variable	Label	Value labels
occ	Occupation	1=Menial 2=BlueCol 3=Craft 4=WhiteCol 5=Prof
white	Race	1 = white 0=nonwhite
ed	Years of education	
exper	Years of work experience	
occ3	Occupation simplified into 3 classes	1=Manual 2=WhiteCol 3=Prof

Multinomial regression: OCC3 = white

mlogit occ3 white

```
Multinomial logistic regression      Number of obs   =      337
                                     LR chi2(2)      =        2.33
                                     Prob > chi2     =       0.3119
Log likelihood = -319.92643          Pseudo R2      =       0.0036
```

occ3	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	

WhiteCol						
white	.808908	.7644154	1.06	0.290	-.6893187	2.307135
_cons	-2.251292	.7433904	-3.03	0.002	-3.70831	-.7942734

Prof						
white	.5465437	.459427	1.19	0.234	-.3539167	1.447004
_cons	-.9985288	.4421411	-2.26	0.024	-1.86511	-.1319481

(occ3==Manual is the base outcome)

now with the option "relative risk" ~ Odds ratio:

mlogit occ3 white,rrr

```
Multinomial logistic regression      Number of obs   =      337
                                     LR chi2(2)      =        2.33
                                     Prob > chi2     =       0.3119
Log likelihood = -319.92643          Pseudo R2      =       0.0036
```

occ3	RRR	Std. Err.	z	P> z	[95% Conf. Interval]	

WhiteCol						
white	2.245455	1.71646	1.06	0.290	.5019179	10.0456

Prof						
white	1.727273	.7935558	1.19	0.234	.7019334	4.250362

(occ3==Manual is the base outcome)

With listcoef you will get an overview of all coefficients and odds ratios for all comparisons

listcoef

mlogit (N=337): Factor Change in the Odds of occ3

Variable: white (sd=.27642268)

Odds comparing Alternative 1 to Alternative 2	b	z	P> z	e^b	e^bStdX
WhiteCol-Prof	0.26236	0.319	0.750	1.3000	1.0752
WhiteCol-Manual	0.80891	1.058	0.290	2.2455	1.2506
Prof -WhiteCol	-0.26236	-0.319	0.750	0.7692	0.9300
Prof -Manual	0.54654	1.190	0.234	1.7273	1.1631
Manual -WhiteCol	-0.80891	-1.058	0.290	0.4453	0.7996
Manual -Prof	-0.54654	-1.190	0.234	0.5789	0.8598

Example to read:

The Odds ratio of a "White Collar job" against a "Manual work job" is 2.246 times higher for white persons than for black people.

Now lets recalculate this Odds ratio with probabilities

With "prtab" we can find the probabilities and recalculate the odds ratio from listcoef above:

For "white" we have Odd for "White collar"/"Manual" = 0.1262 / 0.5340 = 0.2363
 For "non white" we have Odd for "White collar"/"Manual" = 0.0714 / 0.6786 = 0.1052
 The odds ratio 0.2363 / 0.1052 = 2.246

prtab white

mlogit: Predicted probabilities for occ3

Predicted probability of outcome 2 (WhiteCol)

```
-----
      Race | Prediction
-----+-----
NonWhite |    0.0714
  White   |    0.1262
-----
```

Predicted probability of outcome 3 (Prof)

```
-----
      Race | Prediction
-----+-----
NonWhite |    0.2500
  White   |    0.3398
-----
```

Predicted probability of outcome 1 (Manual)

```
-----
      Race | Prediction
-----+-----
NonWhite |    0.6786
  White   |    0.5340
-----
```

```
white
x= .91691395
```

You can also find this probabilities in this example by using the cross table

tab2 white occ3, row

Race	Occupation simplified into 3 classes			Total
	Manual	WhiteCol	Prof	
NonWhite	19 67.86	2 7.14	7 25.00	28 100.00
White	165 53.40	39 12.62	105 33.98	309 100.00
Total	184 54.60	41 12.17	112 33.23	337 100.00

After these demonstrations of how to read and interpret odds ratios it is quite obvious why people prefer ordinary probabilities to interpret and present results.

Long & Freese developed the routines "mlogview" (menu) or "mlogplot" (direct input) to produce graphical results in multinomial logistic regression.

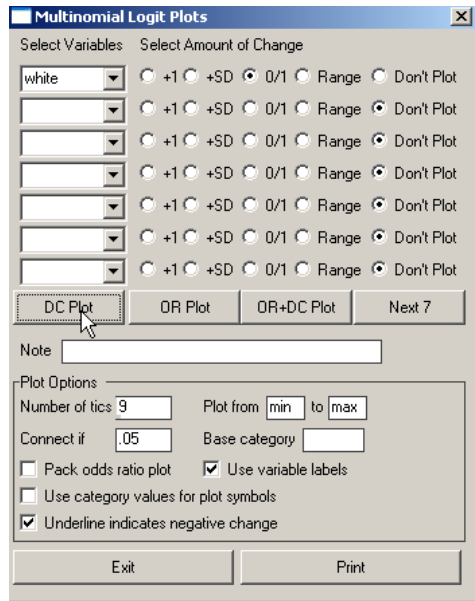
Before you can call this routine you have to call "prchange" so that "mlogview" can use its internal results.

prchange

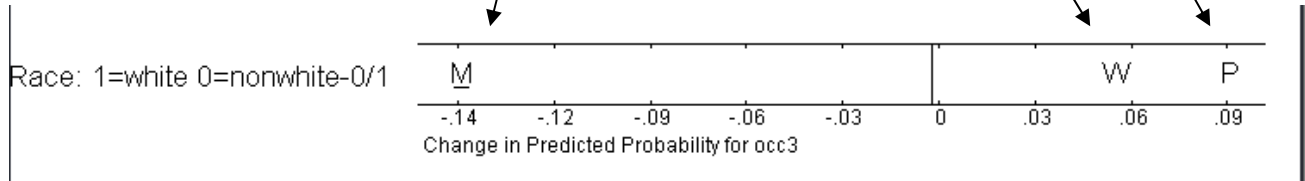
```
mlogit: Changes in Probabilities for occ3
white
      Avg|Chg|   WhiteCol   Prof   Manual
0->1  .09639388 .05478502 .08980581 -.1445908
      WhiteCol   Prof   Manual
Pr(y|x) .12082359 .33246318 .54671323
white
x= .916914
sd(x)= .276423
```

mlogview

Now, we choose DC Plot = Discrete change Plot

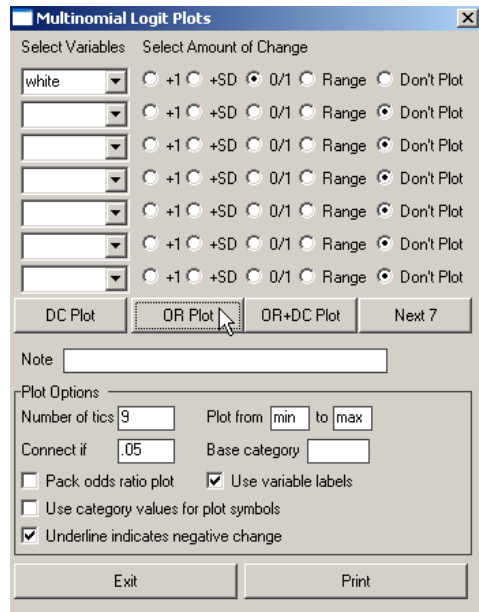


alternative direct: mlogplot white, std(0) p(.05) dc labels sign ntics(9)

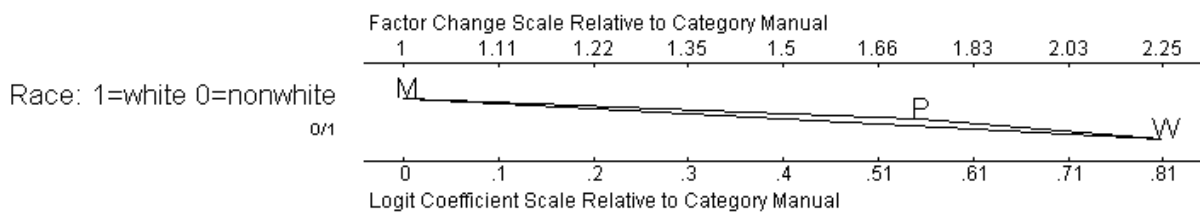


You can also plot the odds ratios:

mlogview



alternative direct: `mlogplot white, std(0) p(.05) or labels sign ntics(9)`



Where

- W = $OR_{whiteCol, Manual}$ (white, nonwhite)
- P = $OR_{prof, Manual}$ (white, nonwhite)
- M = $OR_{Manual, Manual}$ (white, nonwhite)=1

A line between two categories indicates: "**not** a significant difference"

Multinomial regression: OCC3 = white + education + experience

mlogit occ3 white ed exper

```
Multinomial logistic regression      Number of obs   =      337
LR chi2(6)                          =      154.22
Prob > chi2                          =      0.0000
Log likelihood = -243.98352          Pseudo R2       =      0.2401
```

occ3	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]

WhiteCol					
white	1.018823	.7930599	1.28	0.199	-.5355457 2.573192
ed	.3390136	.087539	3.87	0.000	.1674402 .5105869
exper	.0188215	.0128899	1.46	0.144	-.0064423 .0440852
_cons	-7.060006	1.515697	-4.66	0.000	-10.03072 -4.089295

Prof					
white	1.240769	.6196045	2.00	0.045	.0263662 2.455171
ed	.761776	.0844286	9.02	0.000	.596299 .927253
exper	.0194976	.0116698	1.67	0.095	-.0033748 .0423701
_cons	-12.30889	1.44883	-8.50	0.000	-15.14855 -9.469237

(occ3==Manual is the base outcome)

estimates store full

mlogit occ3 white ed

```
Multinomial logistic regression      Number of obs   =      337
LR chi2(4)                          =      150.43
Prob > chi2                          =      0.0000
Log likelihood = -245.87455          Pseudo R2       =      0.2343
```

occ3	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]

WhiteCol					
white	.9570256	.7910705	1.21	0.226	-.593444 2.507495
ed	.3127784	.0866336	3.61	0.000	.1429796 .4825771
_cons	-6.283392	1.421957	-4.42	0.000	-9.070375 -3.496408

Prof					
white	1.191351	.6175922	1.93	0.054	-.0191072 2.40181
ed	.7392936	.0827332	8.94	0.000	.5771395 .9014478
_cons	-11.58107	1.363245	-8.50	0.000	-14.25298 -8.909155

(occ3==Manual is the base outcome)

estimates store short

lrtest full short

```
Likelihood-ratio test              LR chi2(2)     =      3.78
(Assumption: short nested in full) Prob > chi2     =      0.1509
```

Just to check whether STATA is able to detect the full model itself:

lrtest short full

```
Likelihood-ratio test              LR chi2(2)     =      3.78
(Assumption: short nested in full) Prob > chi2     =      0.1509
```


mlogit occ3 white ed exper

```
Multinomial logistic regression            Number of obs    =      337
                                           LR chi2(6)       =     154.22
                                           Prob > chi2      =      0.0000
Log likelihood = -243.98352              Pseudo R2       =      0.2401
```

```
-----+-----
```

occ3	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
WhiteCol						
white	1.018823	.7930599	1.28	0.199	-.5355457 2.573192	
ed	.3390136	.087539	3.87	0.000	.1674402 .5105869	
exper	.0188215	.0128899	1.46	0.144	-.0064423 .0440852	
_cons	-7.060006	1.515697	-4.66	0.000	-10.03072 -4.089295	
Prof						
white	1.240769	.6196045	2.00	0.045	.0263662 2.455171	
ed	.761776	.0844286	9.02	0.000	.596299 .927253	
exper	.0194976	.0116698	1.67	0.095	-.0033748 .0423701	
_cons	-12.30889	1.44883	-8.50	0.000	-15.14855 -9.469237	

```
-----+-----
```

(occ3==Manual is the base outcome)

mlogtest, lr

**** Likelihood-ratio tests for independent variables (N=337)

Ho: All coefficients associated with given variable(s) are 0.

```
-----+-----
```

	chi2	df	P>chi2
white	4.949	2	0.084
ed	150.523	2	0.000
exper	3.782	2	0.151

```
-----+-----
```

This alternative to lrtest is preferred by Long & Freese because it "automatically computes the tests for all variables by making repeated calls to rtest" (p.236).

Later we will see another application of this test: Combining categories

Now let's see the result for listcoef, prtab, prchange and mlogview

listcoef white, help

mlogit (N=337): Factor Change in the Odds of occ3

Variable: white (sd=.27642268)

Odds comparing Alternative 1 to Alternative 2	b	z	P> z	e^b	e^bStdX
WhiteCol-Prof	-0.23433	-0.270	0.787	0.7911	0.9373
WhiteCol-Manual	0.95703	1.210	0.226	2.6039	1.3028
Prof -WhiteCol	0.23433	0.270	0.787	1.2641	1.0669
Prof -Manual	1.19135	1.929	0.054	3.2915	1.3900
Manual -WhiteCol	-0.95703	-1.210	0.226	0.3840	0.7676
Manual -Prof	-1.19135	-1.929	0.054	0.3038	0.7194

 b = raw coefficient
 z = z-score for test of b=0
 P>|z| = p-value for z-test
 e^b = exp(b) = factor change in odds for unit increase in X
 e^bStdX = exp(b*SD of X) = change in odds for SD increase in X

listcoef white, pvalue(0.1)

mlogit (N=337): Factor Change in the Odds of occ3 when P>|z| < 0.10

Variable: white (sd=.27642268)

Odds comparing Alternative 1 to Alternative 2	b	z	P> z	e^b	e^bStdX
Prof -Manual	1.19135	1.929	0.054	3.2915	1.3900
Manual -Prof	-1.19135	-1.929	0.054	0.3038	0.7194

listcoef ed

mlogit (N=337): Factor Change in the Odds of occ3

Variable: ed (sd=2.9464271)

Odds comparing Alternative 1 to Alternative 2	b	z	P> z	e^b	e^bStdX
WhiteCol-Prof	-0.42652	-4.710	0.000	0.6528	0.2846
WhiteCol-Manual	0.31278	3.610	0.000	1.3672	2.5133
Prof -WhiteCol	0.42652	4.710	0.000	1.5319	3.5138
Prof -Manual	0.73929	8.936	0.000	2.0945	8.8311
Manual -WhiteCol	-0.31278	-3.610	0.000	0.7314	0.3979
Manual -Prof	-0.73929	-8.936	0.000	0.4775	0.1132

prtab ed white, outcome(1)

mlogit: Predicted probabilities of outcome 1 (Manual) for occ3

Years of education	Race: 1=white 0=nonwhite	
	NonWhite	White
3	0.9952	0.9875
6	0.9872	0.9668
7	0.9820	0.9534
8	0.9744	0.9339
9	0.9630	0.9050
10	0.9454	0.8614
11	0.9174	0.7960
12	0.8724	0.7010
13	0.8011	0.5739
14	0.6940	0.4257
15	0.5509	0.2822
16	0.3907	0.1683
17	0.2461	0.0924
18	0.1401	0.0480
19	0.0742	0.0240
20	0.0376	0.0118

white ed
 x= .91691395 13.094955

prtab ed white, outcome(2)

mlogit: Predicted probabilities of outcome 2 (WhiteCol) for occ3

Years of education	Race: 1=white 0=nonwhite	
	NonWhite	White
3	0.0047	0.0123
6	0.0120	0.0307
7	0.0164	0.0414
8	0.0222	0.0554
9	0.0300	0.0734
10	0.0403	0.0956
11	0.0534	0.1208
12	0.0695	0.1454
13	0.0872	0.1627
14	0.1033	0.1651
15	0.1122	0.1496
16	0.1088	0.1220
17	0.0936	0.0916
18	0.0729	0.0650
19	0.0528	0.0445
20	0.0365	0.0299

white ed
 x= .91691395 13.094955

prtab ed white, outcome(3)

mlogit: Predicted probabilities of outcome 3 (Prof) for occ3

Years of education	Race: 1=white 0=nonwhite	
	NonWhite	White
3	0.0001	0.0003
6	0.0008	0.0025
7	0.0016	0.0052
8	0.0034	0.0106
9	0.0070	0.0216
10	0.0143	0.0430
11	0.0292	0.0833
12	0.0581	0.1536
13	0.1117	0.2634
14	0.2027	0.4092
15	0.3369	0.5682
16	0.5005	0.7097
17	0.6603	0.8160
18	0.7871	0.8871
19	0.8731	0.9314
20	0.9259	0.9583

white ed
 x= .91691395 13.094955

Now let's plot the probabilities

First, we have to calculate all probabilities we wanted to plot

prgen ed, x(white=1) from(6) to (20) generate(white) ncases(15)
prgen ed, x(white=0) from(6) to (20) generate(nowwhite) ncases(15)

These two commands generate each 7 variables with values

variable	Label
whitex	Years of education ← this is the x axes for the plot
whitep2	pr(WhiteCol)=Pr(2)
whitep3	pr(Prof)=Pr(3)
whitep1	pr(Manual)=Pr(1)
whites2	pr(y<=2)
whites3	pr(y<=3)
whites1	pr(y<=1)
nowhitex	Years of education ← this is the x axes for the plot (same values like the one above)
nowhitp2	pr(WhiteCol)=Pr(2)
nowhitp3	pr(Prof)=Pr(3)
nowhitp1	pr(Manual)=Pr(1)
nowhites2	pr(y<=2)
nowhites3	pr(y<=3)
nowhites1	pr(y<=1)

From these new variables we need

- the variables ending in ...p2 ...p3 ...p1 they contain the probabilities for the three choices
- one of variables ending in ...x they contain the values for the x axes (years of education)

We do not need the new variables ending in ...s2 ...s3 ...s1 because they contain the sum of probabilities (used in ordinal logit models).

First we label the variables ... p1 ...p2 ...p3 so that the legend of our plot is automatically OK

NOTE: Remember (record) the meaning of the variables described in the label !!!

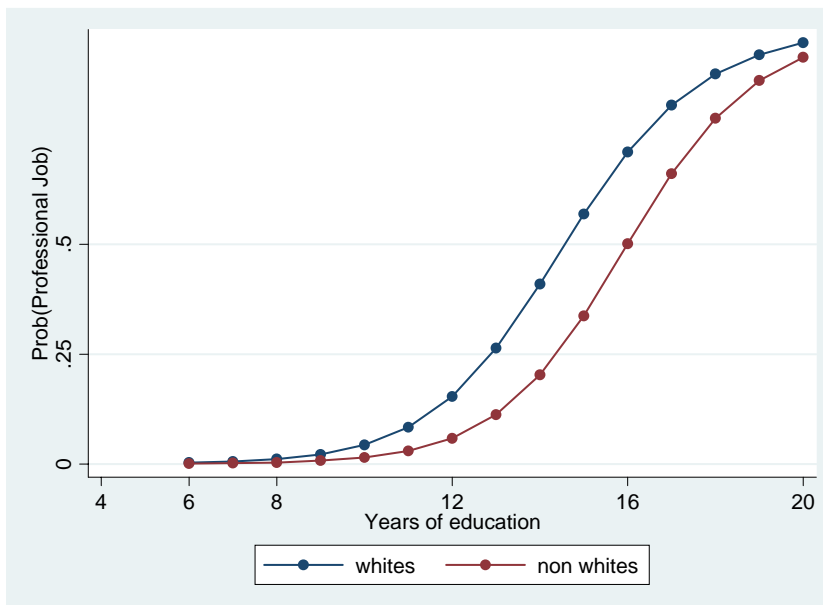
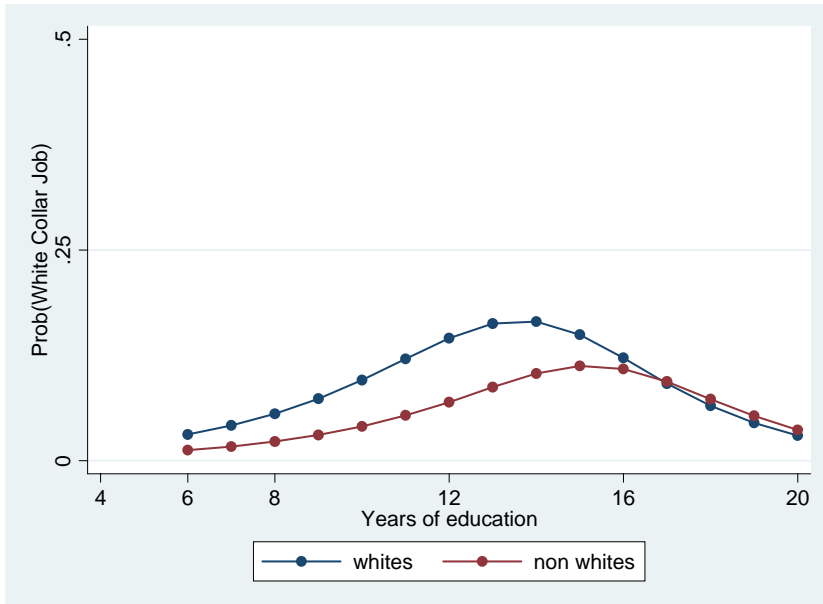
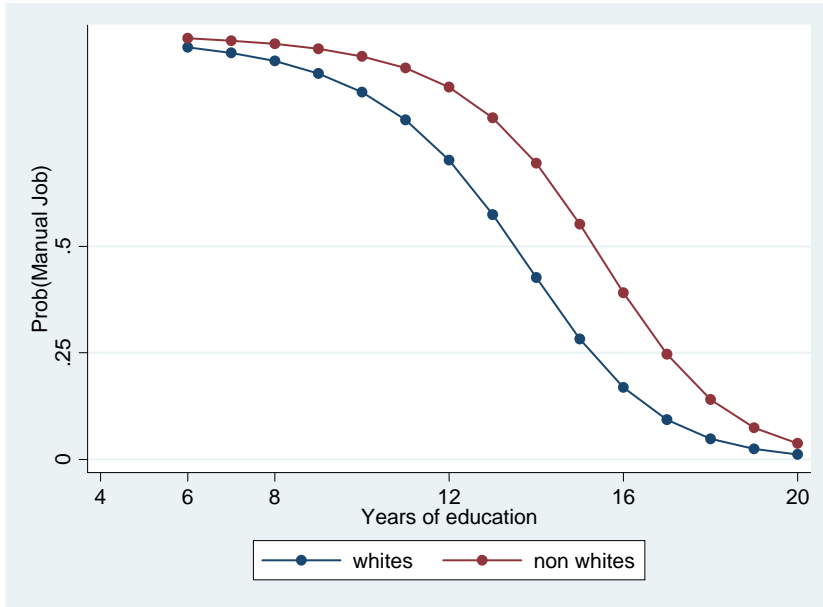
```
label var whitep1 "whites"
label var whitep2 "whites"
label var whitep3 "whites"
label var nowhitp1 "non whites"
label var nowhitp2 "non whites"
label var nowhitp3 "non whites"
```

```
graph twoway connected whitep1 nowhitp1 whitex ,
xtitle("Years of education")
yttitle("Prob(Manual Job)")
ylabel(0 (0.25) 0.5)
xlabel(4 6 8 12 16 20)
```

```
graph twoway connected whitep2 nowhitp2 whitex ,
xtitle("Years of education")
yttitle("Prob(White Collar Job)")
ylabel(0 (0.25) 0.5)
xlabel(4 6 8 12 16 20)
```

```
graph twoway connected whitep3 nowhitp3 whitex ,
xtitle("Years of education")
yttitle("Prob(Professional Job)")
ylabel(0 (0.25) 0.5)
xlabel(4 6 8 12 16 20)
```

graph twoway (see commands above)



Checking the variables on the right side of the model

This can be done via `lrtest` see [above](#)
 or easier via `mlogtest, lr` see [above](#)

Checking the categories (choices) on the left side of the model (indistinguishable)

To test, whether we can use a model with combined alternatives (choices) for the left side variable (dependent) Which is a test that alternatives are indistinguishable in respect to variables in the model.

```
mlogit occ white ed exper, baseoutcome(5) nolog /* base category is 5 (Prof) */
```

1. Wald test to test all J*(J-1) pairs of categories "mlogtest, combine"

```
mlogtest, combine
```

```
**** Wald tests for combining alternatives (N=337)

Ho: All coefficients except intercepts associated with a given pair
of alternatives are 0 (i.e., alternatives can be combined).
Alternatives tested|      chi2   df   P>chi2
-----+-----
Menial- BlueCol   |      3.994    3   0.262
Menial-  Craft   |      3.203    3   0.361
Menial-WhiteCol  |     11.951    3   0.008
Menial-  Prof   |     48.190    3   0.000
BlueCol-  Craft   |      8.441    3   0.038
BlueCol-WhiteCol |     20.055    3   0.000
BlueCol-  Prof   |     76.393    3   0.000
  Craft-WhiteCol  |      8.892    3   0.031
  Craft-  Prof   |     60.583    3   0.000
WhiteCol-  Prof   |     22.203    3   0.000
-----+-----
```

2. Wald test to test a specific alternative against the base category "test [category]"

```
test [Menial]
```

```
( 1) [Menial]white = 0
( 2) [Menial]ed = 0
( 3) [Menial]exper = 0

      chi2( 3) =      48.19
      Prob > chi2 =      0.0000
Is "Menial" undistinguishable from the base category "Prof" (the same as line Menial-Prof above)
```

3. Wald test to test any alternative against any other "test [category1=category2]"

```
test [Menial=Craft]
```

```
( 1) [Menial]white - [Craft]white = 0
( 2) [Menial]ed - [Craft]ed = 0
( 3) [Menial]exper - [Craft]exper = 0

      chi2( 3) =      3.20
      Prob > chi2 =      0.3614
```

4. Likelihood-ratio test (LR) to test all J*(J-1) pairs of categories "mlogtest, lrcomb"

```
mlogtest, lrcomb
```

```
**** LR tests for combining alternatives (N=337)

Ho: All coefficients except intercepts associated with a given pair
of alternatives are 0 (i.e., alternatives can be collapsed).
Alternatives tested|      chi2   df   P>chi2
-----+-----
Menial- BlueCol   |      4.095    3   0.251
Menial-  Craft   |      3.376    3   0.337
Menial-WhiteCol  |     13.223    3   0.004
Menial-  Prof   |     64.607    3   0.000
BlueCol-  Craft   |      9.176    3   0.027
BlueCol-WhiteCol |     22.803    3   0.000
BlueCol-  Prof   |    125.699    3   0.000
  Craft-WhiteCol  |      9.992    3   0.019
  Craft-  Prof   |     95.889    3   0.000
WhiteCol-  Prof   |     26.736    3   0.000
-----+-----
```

Assumption: Independence of irrelevant alternatives (I I A)

Not-binary logit models do have additional assumptions³.

In case of multinomial logit model (MNL) and conditional logit models it is the assumption "independence of irrelevant alternatives = IIA).

The idea of that assumption is, that the odds for a decision are not depending on other categories.
E.g. one could think of a situation where the odd for a decision would become smaller and smaller the more alternatives are possible - that should be excluded by the IIA assumption.

Unfortunately there seems to be no reliable test available to test whether the assumption is violated.

Long and Freese discuss the often used tests Hausman-McFadden and Small-Hsui and come to the conclusion "we do not encourage their use." (p. 244)⁴

They suggest:

"It appears that the best advice regarding IIA goes back to an early statement by McFadden (1973), who wrote that multinomial and conditional logit models should be used only in cases where the alternatives can 'plausibly be assumed to be distinct and weighted independently in the eyes of the decision maker'."

³ In case of ordered models it is the "parallel regression assumption" or the "proportional odds assumption"

⁴ Whereas two tests (an approximate LR-Test <omodel> and a Wald test by Brant <brant>) are available to test the parallel regression assumption in ordered models (p.199).